Causal Inference in Population Trends: Searching for Demographic Anomalies in Big Data

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Abstract

The proliferation of big data, wider access to advanced computing platforms, and the development of powerful statistical algorithms can uncover hidden anomalies in social data, previously dismissed as noise. Here, we combine causal inference techniques and abductive reasoning to identify fertility and mortality anomalies on twenty years of complete demographic data in the United States. We uncover real, “hidden” baby booms/busts and mortality spikes/dips, distinguishable from regular trend variations. We identify more than 22 and 156 fertility and mortality anomalies, totaling more than 200k and 600k anomalous births and deaths, respectively. Notable detectable mortality anomalies include the September 11 2001 terrorist attack in New York and the emergence and acceleration of the opioid epidemic in New Hampshire. Notable fertility anomalies include the “missing births” in Louisiana after Hurricane Katrina and the reduction in fertility behavior after the September 2008 stock market crash in Connecticut, amongst others. The combined causal inference and abductive reasoning approach can be readily adapted to find other, undiscovered social phenomena or to evaluate the efficacy of important public policies.

Keywords: mortality, fertility, causal inference, abductive reasoning, big data

1 Introduction

The search for causality is one of the primary motivations of social scientists (Smith 2011, Pearl et al. 2009). Causal inference, the identification of causal mechanisms, attempts to uncover the underlying mechanisms that result in changes in a phenomenon and has a long history as an approach to inquiry in the social sciences (Grimmer 2015). Often, the identification of a casual mechanism is accomplished via a randomized control trial (RCT) or expert knowledge applied to a natural experiment (Salganik 2019).

Unfortunately, both RCTs and natural experiments have drawbacks. RCTs are often prohibitively expensive for many scientists, precluding their widespread adoption in the social sciences (West et al. 2008). RCTs are also sometimes impossible to deploy in the “real world.” For instance, an RCT to study country-level austerity measures would be difficult, if not impossible, to implement. RCTs to study fertility would be immoral. Natural experiments are a cheaper alternative but provide less certainty about causation, fewer opportunities to replicate work, and require expert knowledge of the natural experiment to be properly employed (Pearl & Mackenzie 2018).

Additionally, both RCTs and natural experiments are based on traditional hypothesis-based testing, the deductive scientific paradigm where a question is first posed and scientists generate or find data to answer the question. Such approaches, while extensively utilized and time-tested, increase the likelihood that a researcher will miss important phenomena that might go unnoticed because it falls outside the hypothesis, or the study could suffer from extensive p-hacking (Head et al. 2015, Nuzzo 2014, Ruggles 2014). At the same time, purely inductive data mining can result in false assertions or spurious findings, especially in the hands of those who are not field experts (Yanai & Lercher 2020). Abductive reasoning, sometimes described as “inferring cause from effect” (Crowder & Carbone 2017), moves the scientific process from the inductive to the deductive, and sometimes back and forth, to reach conclusions (Bryant & Raja 2014). Physical scientists use abductive reasoning more commonly than social scientists. The discovery of Neptune in 1846 based on unexplained perturbations of the orbit of Uranus (Popper 2005), the discovery of galactic migration of planets (Gomes et al. 2005), and the current generation of gravitational wave detectors such as the Laser Interferometer Gravitational-Wave Observatory (LIGO) (Harry et al. 2010)

demonstrate the power of the abductive approach in the physical sciences.

In a data-rich environment, abductive reasoning can be more easily applied to first find effects and then infer their cause ex-post. The modern proliferation of data, advances in high performance (“super”) computing, and the development of powerful statistical algorithms provide an environment ripe for wedding causal inference with abductive reasoning (Van Der Aalst 2016, Zikopoulos & Eaton 2011), with the potential for researchers to find hidden or understudied social phenomena (Bohon 2018). While the availability of big data and high performance computing allows novel exploration of data through causal inference (Bohon 2018, Brodersen et al. 2014, Shiffrin 2016), relatively few studies utilize causal inference techniques in the study of demographic phenomena. However, understanding social phenomena using these advances reveals important insights into society (Angrist et al. 1989, Mas & Moretti 2009) and allows us to better monitor population trends (Nobles & Seltzer 2019, Torche & Shwed 2015). The rich demographic data available in the United States make the potential revelation of interesting and important demographic phenomena not only possible, but extremely plausible – even if identification of the phenomena occurs without identifying the underlying cause.

In this paper, we combine a modern statistical outlier detection algorithm (Chen & Liu 1993), abductive reasoning, and large demographic data sets on more than 20 years of mortality and fertility trends to identify anomalous demographic behavior. In essence, we identify *effects without knowledge of the cause*. We ask one fundamental question regarding demographic anomalies: What are the hidden baby booms/busts and mortality spikes/dips in the United States over the last twenty years? We do not necessarily know the causes of these anomalies and suggest causes for select anomalies. Even without knowledge of the causes, simply identification of anomalies allows scientists with more detailed knowledge of local population dynamics, state-level policy making, or macro-economics to explain these phenomena post-hoc. From explanations of phenomena that may have previously gone unnoticed, social scientists may be able to better forecast populations and provide policy solutions for impending problems such as climate change.

2 Abductive Reasoning and Causal Inference

Scientific inquiry falls under one of three modes of logical reasoning: deduction, induction, and abduction. Deductive reasoning starts with a theory and makes a prediction of what the observations should be if that theory is correct. The classic example of deductive reasoning essentially goes (a) “all men are mortal,” (b) “Socrates is a man,” (c) “therefore, Socrates is mortal.” Deductive reasoning has a long history in the sciences and most hypothesis testing in the social sciences follows this logical reasoning. For example, John Snow used deductive reasoning to stop the 1854 London cholera outbreak. He hypothesized that cholera was waterborne and deduced that a bad water pump was responsible for the outbreak.

Inductive reasoning is directionally opposite of deductive reasoning. Rather than starting with a theory and making a prediction, induction starts with an observation and makes a theory based on the likeliest reason. For example, the Hispanic Mortality Paradox first appeared in the scientific literature in the mid-1980s (Markides & Coreil 1986) as an observation that Hispanics in the United States had better health outcomes than expected given their socioeconomic status. Based on this observation, numerous theories arose to explain this paradox.

Both deductive and inductive reasoning are well established in science but are not without their shortcomings. In both approaches, it is impossible, or at least improbable, for the premise to be true and the conclusion false. Furthermore, if a hypothesis is rejected under a deductive or inductive approach, there is no alternative explanation – there is either the presence or absence of support.

Abductive reasoning is a third, less well-known way that reasons the best explanation given the observations (Walton 2014). It is most concerned with generating and evaluating explanatory hypotheses (Crowder & Carbone 2017) by looking for a pattern and suggesting a hypothesis (Fann 2012, Pierce 1878). Unlike the more rigid deductive and inductive reasoning, abductive reasoning allows the scientist to move between deduction and induction and fits well during exploratory data analysis (Yu 1994, Haig 2015, Tukey 1977). During exploratory data analysis, abductive reasoning helps find a pattern and suggest a hypothesis, deduction helps refine and build a testable hypothesis, and induction is the empirical

support (Yu 1994).

Abductive reasoning and exploratory data analysis are best employed in a data rich environment – the very data environment social scientists and demographers find themselves in the present moment. In a data rich environment, a hypothesis is a hidden liability (Yanai & Lercher 2020) because it focuses our attention on only that which we seek. To demonstrate this, ¹. Focusing on specific hypotheses often prevents us from finding other interesting phenomena. After all, it was the publication of the Hispanic paradox itself that has generated more than thirty years of devoted scholarship. At the same time, when paired with causal inference techniques, abductive reasoning provides an empirical platform for the discovery of potentially interesting social phenomena; it provides an important tool for sifting large amounts of data to find interesting results without the blinders of hypothesis testing.

3 Materials and Methods

3.1 Method

We use a statistical time series outlier detection algorithm (Chen & Liu 1993), implemented in the R programming language (R Core Team 2019) via the *tsoutliers* package (López-de-Lacalle 2019). This algorithm iteratively uses ARIMA models to 1) identify potential outliers or anomalies and 2) refit the ARIMA with the outliers removed to produce a counter-factual time series. Here we briefly summarize and describe the method.

Often, the behavior of a time series can be described and summarized in ARIMA models. If a series of values, y_t^* , is subject to m interventions or outliers at time points $t_1, t_2, \dots, t_m$ with weights ω then y_t^* can be defined as

$$y_t^* = \sum_{j=1}^m \omega_j L_j(B) I_t(t_j) + \frac{\theta(B)}{\phi(B)\alpha(B)} \alpha_t \quad (1)$$

¹Yanai and Lercher (2020) ask if focusing on a specific hypothesis prevents new discovery. By way of illustration, they created a simple dataset that if plotted would reveal an image of a gorilla waving. Students instructed to find support for specific hypotheses were less likely to discover the gorilla compared to those who were simply asked “What do you conclude from the dataset?”

Where $I_t(t_j)$ is an indicator variable with a value of 1 at observation t_j and where the j th outlier arises,

$\phi(B)$ is an autoregressive polynomial with all roots outside the unit circle,

$\theta(B)$ is a moving average polynomial with all roots outside the unit circle,

and $\alpha(B)$ is an autoregressive polynomial with all roots on the unit circle.

We examine three types of outliers at time point t_m :

1. additive outliers (AO), defined as $L_j(B) = 1$;
2. level shift outliers (LS), defined as $L_j(B) = 1/(1 - B)$; and
3. temporary change outliers (TC), defined as $L_j(B) = 1/(1 - \delta B)$ where δ is equal to 0.7.

Colloquially, additive outliers arise when a single event causes the time series to unexpectedly increase/decrease for a single time period; level shift outliers arise when an event causes the time series to unexpectedly increase/decrease for multiple time periods; and temporary change outliers arise when an event causes the time series to unexpectedly increase/decrease with lingering effects that decay over multiple time periods.

An outlier is detected with the estimated residuals using a regression equation

$$\pi(B)y_t^* \equiv \hat{e} = \sum_{j=1}^m \omega_j \pi(B) L_j(B) I_t(t_j) + \alpha_t \quad (2)$$

where $\pi(B) = \sum_{i=0}^{inf} \pi_i B^i$.

Equations 1 and 2 allow for an automatic detection of anomalies iterated over a two-stage process.

In stage 1, anomalies are located. First, an ARIMA model is fit to the time series using the **forecast** package in R (Hyndman et al. 2019, Hyndman & Khandakar 2008) where the best performing ARIMA model is selected based on the Bayesian information criterion (BIC). Next, the residuals from the forecast are checked for their significance using equation 2 where only anomalies above a critical t -static are considered “true” anomalies ($|\tau| \geq 3.5$). We choose this critical value based on Chen and Liu’s (1993) recommendation to minimize Type I or false-positive anomalies. Finally, two additional rules are implemented: If multiple anomalies are detected at the same time point, only the most significant anomaly is selected and if anomalies of the same type at consecutive time periods are detected, only the anomaly with highest t -statistic is selected.

In stage 2, anomalies are removed from the time series and a new ARIMA model is chosen and fit. The selection of the initial ARIMA model could have been affected by the presence of the anomalies, making some anomalies spuriously identified. To correct for this, a new ARIMA model is fit accounting for additional regression effects in equation 1 from the list of candidate anomalies identified in stage 1, effectively removing the anomalies from the time series. Each anomaly is then reassessed under the new model and those anomalies that are no longer significant are removed.

These two stages are then iterated until no additional anomalies are detected.

3.2 Robustness Check

To determine the false-positive (Type I) and false-negative (Type II) error rates under a critical value of 3.5, we ran two simulations. For the false-positive error we simulated 1000 time series with 180 time steps without introducing any anomalies and ran the `tsoutliers` algorithm on them. Any anomalies we detect in these time series are spurious anomalies. For the false-negative error, we simulated 1000 time series with 180 time steps but introduced a single anomaly of the same magnitude at the same time point in each time series. If the `tsoutliers` algorithm fails to detect the anomaly, it would be considered a false-negative error.

In our sans-anomaly time series, we detect 208 anomalies on the 180,000 candidate time steps for a false-positive error rate of 0.116%. In our time series containing a single anomaly, we detect 817 anomalies out of the 1000 candidate time series, for a false-negative error rate of 18.3%. This provides a statistical power of 0.817. These results give us confidence that the anomalies we detect in this paper are true anomalies.

3.3 Toy example

Figure 1 shows a toy example for anomaly detection in a time series using the classic example of the exchange rate between the Euro and the Swiss Franc (CHF). On January 15 2015, the Swiss National Bank removed the currency peg of 1.20 francs per Euro, exposing the Franc to the volatility of the currency market. The Franc immediately began trading a reduced rate compared to the Euro. What would have been the Euro/CHF exchange

Figure 1: **“Toy” example of anomaly detection using the exchange rate between the Euro and the Swiss Franc.** The Swiss Franc was pegged to the Euro at 1.2 until January 15 2015 when the peg was removed and the currency was allowed to trade more freely. This is a very strong anomaly in the time series ($t\text{-stat} = -144.07$) detectable using statistical methods and visually. The counterfactual estimate with the outlier removed would simply have kept the currency exchange at 1.2 ($\hat{y}$).

rate had the Swiss National Bank *not* removed the peg? A simple counter-factual estimate would be to keep the exchange rate at 1.20 ($\hat{y}$, sometimes called a “synthetic control” (Abadie et al. 2010)).

In the Swiss Franc example, we have knowledge of the Swiss Bank’s activities after the fact or ex-post-facto to create the counter-factual time series $\hat{y}$ of 1.20. But is this anomaly detectable without knowledge of the Swiss Bank’s activities? In other words, can we detect the reduction of the exchange rate on January 15 using only the time series? Absolutely. the `tsoutliers` package identifies January 15 2015 as an extremely strong level-shift outlier (t -stat = -144.07). The real world is hardly this simplified where a direct intervention is known and is testable.

3.4 Data

We search for demographic anomalies using the Center for Disease Control and Prevention’s online WONDER monthly fertility (2003-2017) and mortality databases (1999-2016) for all fifty states and the District of Columbia (United States Department of Health and Human Services (US DHHS), Centers for Disease Control and Prevention (CDC) 2018, Centers for Disease Control and Prevention, National Center for Health Statistics. 2019). These data sets contain every birth and death record in the United States over the time periods of interest, representing the universe of both mortality and fertility data in the US. These data are considered the “gold standard” of data collections (Mahapatra et al. 2007) and have been considered “complete” since 1968 (Hetzl 2016). We search over each state equivalent’s ($n=51$) mortality ($n=228$) and fertility ($n=180$) monthly time series for a total of 20,808 state-months of data.

3.5 Replication Files

All data and code necessary to reproduce the reported results are licensed under the CC-BY-4.0 license and are publicly available as https://osf.io/yc3r4/?view_only=3ce3ca267c7047e0942940ebd08846d0. The analyses were performed in *R* (R Core Team 2019).

4 Results

We detect numerous anomalous mortality and fertility events at the US state-level since 1999. A full listing of these anomalies can be found in the **Supplementary Materials**. We begin by highlighting significant fertility and mortality anomalies across all three types of outliers (Additive Outliers, Level Shift Outliers, and Temporary Change outliers) with plausible explanations. We then highlight two strong anomalies that bely explanation. Finally, we conclude with a summary of the anomalies we detect.

4.1 Fertility

We begin with a clear temporary change (TC) outlier for fertility in Louisiana (**Figure 2a**). Here we detect two TC outliers, nearly back-to-back in 2005 (2005:08 $t = -6.002$; and 2005:10 $t = -4.997$), coinciding with the destruction of Hurricane Katrina from the same year. The migration associated with Hurricane Katrina has garnered most of the attention of social scientists (Fussell et al. 2014, Hori et al. 2009), but the hurricane had a very clear impact on fertility behaviors too, creating a mini “baby bust” in Louisiana with the departure of many people in their childbearing years. We estimate 4,411 fewer births in Louisiana compared to the counter-factual, likely attributable to the Hurricane.

We also highlight a level shift (LS) outlier for fertility in Connecticut. Here we detect a strong ($t = -3.71$ for -285 births/month) level shift toward lower fertility beginning in August 2009 (2009:08) that continues through the end of the period (**Figure 2b**). This LS toward lower fertility reduced the number of births by 28,752 or 9%) lower than the counter-factual time series. The stock market crash of 2008, where the Dow Jones Industrial Average had the largest single-day loss up to that point, occurred just 11 months before we detect a shift toward lower fertility. It is possible that the trend toward lower fertility in August 2009 and the stock market crash in September 2008 are linked, although, as stated previously, we are speculating only to illustrate how causal inference could work in an abductive research agenda. A potentially interesting future line of research would be to link fertility anomalies to the kinds of economic shocks that may lead to out-migration of those in their childbearing years.

a **b**

Figure 2: **Anomaly detection in Louisiana (a) and Connecticut (b) state fertility, 2003-2018.** The top part of each panel contains the original time series (light gray), the corrected, counter-factual time series in the absence of anomalies (blue), and the red dots correspond to the onset of detected anomalies. The bottom part of each panel contains the magnitude and type of the outlier in red. In (a), we detect two outliers, back-to-back, likely resulting from Hurricane Katrina in August and October 2005, representing a decrease of more than 4,400 births due to the hurricane (t -statistics= -6.002 and -4.997). In (b), we detect one outlier, a level shift outlier (LS) in August 2009. We believe this reduction is attributable to the stock market crash 11 months earlier (t -statistic = -3.71).

4.2 Mortality

For mortality, we will use New York State as an example (**Figure 3a**). We identify seven anomalies in the mortality time series for New York, all with t -statistics in excess of 3.91, making these anomalies unlikely to be due to chance. The algorithm correctly identifies September 2001 as an additive outlier (2001:09 $t= 6.396$) where there were 1,628 more deaths in that month than anticipated. This mortality event is likely caused by the September 11 terrorist attack on the World Trade Center that immediately killed 2,606 people and the detection of this mortality event provides confidence in our detection of other anomalies.

In **Figure 3a**, notice the strong level shift (LS) that occurs in February 2004 (2004:02 $t= -5.594$) which prevented 869 deaths per month. This shift totals more than 144,000 averted deaths compared to the counter-factual time series and is the single largest mortality protective anomaly among all states. This translates to 7% fewer deaths than expected over the time period. What is driving this mortality protection? What policies did NY put into place that might have contributed to this considerable mortality reduction? What environmental or economic conditions may have changed? These are the kinds of questions that arise from our analyses. The purpose of our paper is not to answer these questions, but our findings underscore the need for more research that takes an abductive approach. By identifying anomalies through an inductive process, researchers can then look for underlying causes. Once those causes are identified, researchers can then use a deductive process to see if such events predict other (or future) anomalies.

To see the potential for combining abductive reasoning with causal inference, contrast the mortality protection in New York with the enhanced mortality in New Hampshire (**Figure 3b**). In New Hampshire we detect two significant level shifts (LS) in the monthly mortality data, first in April 2010 and again in November 2014 (2010:04 $t= 3.51$; 2014:11 $t= 6.68$). These anomalies suggest New Hampshire experienced 8,159 more deaths (+9% more than expected) in a seven-year period beginning in early 2010. These are events not experienced by neighboring states during the same time period, and this is the single largest percentage mortality increase/decrease we detected among all states. Not coincidentally, NH has the second highest opioid-related mortality in the US (Beetham et al. 2019) and it is likely that we detect this epidemic in our results. Isolating these anomalies and testing

them against opioid sales data might yield intriguing results.

4.3 Interesting Anomalous Fertility/Mortality events

In the examples above, we highlighted four fertility/mortality anomalies with plausible explanations. In the case of New York and New Hampshire, the mortality anomalies have plausible explanations. It seems likely that New York’s AO anomaly in September 2001 is caused by the 9/11 tragedy and the rise in New Hampshire’s mortality starting in 2010 could be linked to the opioid epidemic. Similarly, Louisiana’s TC anomalies seem linked to Hurricanes Katrina and Rita while the LS anomaly in Connecticut’s fertility appears linked to the Great Recession. However, we detect numerous other demographic anomalies in other states, on the causes of which we will not speculate. **Figure 4** shows two such unexplained anomalies.

In **Figure 4a** we identify a single additive anomaly in Hawaiian fertility in May 2014. This is a strong anomaly with a t -statistic of 4.96, 15% above the counter-factual time series. This single, anomalous month is also the second highest monthly births in the time series. We have no plausible explanation for this anomaly. We do not believe this is simply a data error as the other extreme values, September 2008 and February 2005 with the highest and lowest recorded fertility respectively, were not identified as anomalous events. Even if we were to assume this anomaly resulted from data entry error, it remains an *unaltered* data error in the Hawaiian monthly fertility data.

This is contrast to **Figure 4b**, where we identify a strange mortality reduction in Ohio (t -stat: -4.27). This LS is more than 1,157 deaths per month less than the counter-factual time series, suggesting Ohio had nearly 40,000 fewer deaths since February 2015 than expected. This is the single-largest LS among all states. We could not identify the potential policies Ohio might have put into place or the events that occurred to provide such a strong mortality protection.

4.4 Overall Anomalies

Table 1 reports the overall number of anomalies we detect of each type for births and deaths and some summary statistics across all anomaly types and **Figure 5** maps these

Figure 3: **Anomaly Detection for New York (a) and New Hampshire (b) state mortality, 1999-2016.** In New York (a), we detect additive outliers (AO) in January 2000, September 2001, January 2005, and January 2013; temporary change (TC) outliers in February 2007 and January 2015; and a level shift (LS) starting in February 2004. In New Hampshire (b), we detect two outliers, both level shift outliers (LS) in April 2010 and again in November 2014. These anomalies suggest New York experienced a significant mortality event in September 2001 and New Hampshire experienced approximately 9,700 more deaths than expected since 2010 or 14% more deaths in the state over just seven years.

a

b

Figure 4: **Anomaly detection in Hawaii state fertility (a) and Ohio state mortality (b).** Here we detect one outlier, an additive outlier (AO) in May 2014 in Hawaii (a), representing a large, unexplainable 14% increase in expected births in that month. We also detect a strong level shift (LS) reduction (b), representing a large, unexplainable reduction of nearly 40,000 deaths in Ohio.

Table 1: **Summary by Demographic Component.** Here we can see there 22 Fertility and 156 Mortality anomalies among the state-level time series, totalling more than 200k anomalous births and 600k anomalous deaths.

Component	Anomalies	y	$\hat{y}$	$y - \hat{y}$	$ y - \hat{y} $	% of Total
Fertility	22	25.48M	25.68M	-202,217	226,476	0.889
Mortality	156	43.57M	43.52M	48,535	627,364	1.440

results.

We find considerably more mortality anomalies ($n = 156$) than fertility anomalies ($n = 22$). Given that anomalies are always the product of events (Song et al. 2018), finding more mortality than fertility anomalies is not surprising. Mortality is likely to spike in response to a catastrophic event (like an earthquake or terrorist attack) or due to a disease outbreak while the effects of a catastrophic event on fertility is less predictable. One reason for this is that fertility is linked to human decision-making more directly than mortality (Stein et al. 2014) which results in more varied outcomes, so, for example, researchers find that catastrophic weather events can increase childbearing among those who already have a child, but not among the childless (Evans et al. 2010). Some events, like the 1995 Oklahoma City bombings, resulted in both fertility and mortality changes. However, changes in fertility manifest over a longer time horizon after an event, and the relationship between a fertility-inducing event and behavior change is less strong than the relationship between a mortality-inducing event (such as a terrorist incident) and death (Rodgers et al. 2005). Of course, not all events that impact fertility and mortality are catastrophic. Researchers have documented that more commonplace events such as massive layoffs (Venkataramani et al. 2019) or policy changes (Livingston et al. 2018) can also create anomalies in mortality patterns, but we have not found such evidence for fertility.

Consistent with more anomalies across the time series, we find more anomalous deaths (627,364) than anomalous births (226,476). Fertility anomalies overwhelmingly tend toward lower fertility, with only a few fertility anomalies yielding more births. This is consistent with research findings that link severe weather events to significantly lower fertility (Evans

et al. 2010). Conversely, mortality anomalies tend to be more evenly split between positive and negative anomalies, but tend toward more deaths rather than fewer deaths. Again, this is not surprising, as the kinds of events that increase death (e.g., plant closings are associated with increased opioid death (Venkataramani et al. 2019)) are more common than the events that decrease death (e.g., cannabis legalization is associated with decreased opioid deaths (Livingston et al. 2018)). These anomalous deaths and births account for 1.44% and 0.889% over the entire time series, respectively.

As **Figure 5** shows, three states exhibited neither mortality nor fertility anomalies: Alaska, North Dakota, and South Dakota. Another nine states exhibited just a single anomaly (Delaware, District of Colombia, Idaho, Illinois, Kansas, Nevada, New Mexico, Utah, and Wyoming). Both New York and Massachusetts exhibited the most anomalies with nine.

5 Conclusion

Data scientists frequently claim that the big data revolution is a turning point in scientific discovery that will allow us to solve some of the world’s most pressing problems (Grimmer 2015). Social scientists are skeptical of such claims, because they better understand the complexities of the social world and know from experience that data alone are not enough to solve social problems (Bohon 2018, Grimmer 2015). Nonetheless, the increased availability of data and (more importantly, we argue) the development of advanced techniques for analyzing these data will enable important discovery (Monroe et al. 2015).

One technique that population scientists underutilize is causal inference. Causal inference, in the simplest terms, is the discovery of effects in search of a cause using big data and advanced computing algorithms (Imai et al. 2008). This inductive approach is uncommon in quantitative social science where hypothesis testing is expected, and approaches are largely deductive. However, common statistical hypothesis testing is impractical with big data, as significant p values are guaranteed, and such approaches do not allow us to uncover all the information that big data has to offer (Monroe et al. 2015). Here, we call for an abductive approach, where causal inference algorithms are applied to high quality data to uncover irregularities that are unlikely to be attributable to expected variations

Figure 5: **Summary of Anomalies by type, demographic component, and State.**
We observe much fewer fertility anomalies than mortality anomalies.

in trends (or noise) as a first step to then developing testable hypotheses about causes. Abductive approaches allow researchers to move from the inductive to the deductive and sometimes work back and forth in aid of scientific discovery.

In this paper, we show how the `tsoutlier` package in R can be implemented to conduct statistical time series outlier detection, an inductive approach that aids in the creation of deductive reasoning. This algorithm is one of many causal inference approaches that are freely and commercially available (see Brodersen et al. (2014)). In our initial work, we experimented with other approaches, such as Google’s CausalImpact algorithm, which uncovered the same patterns we briefly discuss in this paper. By making more use of causal inference techniques and abductive modeling approaches, we argue that social scientists will be able to better understand how events or policy implementation can impact important outcomes. For example, we could uncover—on a wide scale—how gun control policies may reduce or increase injuries from shootings or how marijuana legalization might impact opioid deaths. We could also uncover how increases in extreme weather events impact a range of behaviors such as home sales, bottled water purchases, and even fertility.

In our demonstration, we show how the application of causal inference to state-level time series fertility and mortality data uncovers three types of demographic anomalies: those that occur and disappear quickly, those that occur and decay over time, and those that occur and remain. Uncovering these anomalies in and of themselves is important. For example, the algorithm we deploy clearly shows the fertility effects of Hurricanes Katrina and Rita as well as the mortality effects of the World Trade Center collapse on September 11, 2001. The ability to identify and differentiate types of anomalies is even more important, as we can potentially see how some policies might impact outcomes permanently and some might have an effect that is short-lived. Differentiating these types gives us greater insight into short- and long-term solutions to social problems. We urge social scientists to begin to use causal inference algorithms and other big data techniques, and we hope that this demonstration will illustrate their usefulness to the social science enterprise.

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Supplementary Material

Matt Hauer

9/16/2020

Table 1: **All Demographic Anomalies detected.**

state	STATEID	component	time	type	tstat	coefhat
AL	01	Fertility	2008:11	AO	-3.63526	-476.0911
AL	01	Mortality	2000:01	AO	5.12804	486.7737
AL	01	Mortality	2000:03	TC	-4.55360	-470.9793
AL	01	Mortality	2001:12	AO	-3.62920	-333.0932
AL	01	Mortality	2003:12	AO	7.48327	660.4668
AL	01	Mortality	2008:02	TC	4.45801	410.0415
AL	01	Mortality	2013:01	AO	4.08582	363.8702
AZ	04	Fertility	2005:10	AO	-3.93621	-580.0972
AZ	04	Fertility	2008:03	LS	-3.68783	-457.7238
AZ	04	Mortality	2003:12	AO	4.18375	454.8685
AZ	04	Mortality	2005:12	TC	4.88077	654.6778
AZ	04	Mortality	2016:02	LS	4.63373	555.4859
AR	05	Mortality	2000:01	AO	4.28328	396.4951
AR	05	Mortality	2003:12	AO	4.04010	287.6263
AR	05	Mortality	2008:02	AO	5.08878	382.4616
AR	05	Mortality	2013:01	AO	5.54341	409.2410
AR	05	Mortality	2014:01	TC	3.52244	292.0337
AR	05	Mortality	2017:12	AO	3.95242	359.8083
CA	06	Fertility	2016:08	TC	3.66555	2203.5513
CA	06	Mortality	2003:12	AO	4.39587	1987.4043
CA	06	Mortality	2005:12	AO	4.12364	1807.0251
CA	06	Mortality	2008:02	TC	4.94494	2602.4935
CA	06	Mortality	2017:12	AO	3.72352	2151.3643
CO	08	Mortality	2013:01	AO	3.53049	246.5604
CO	08	Mortality	2015:02	TC	-6.10508	-578.7003
CT	09	Fertility	2009:08	LS	-3.71363	-284.6775
CT	09	Mortality	2003:12	AO	4.04637	296.2717
CT	09	Mortality	2013:01	AO	3.74637	279.6102
CT	09	Mortality	2014:02	LS	5.51927	139.8958
CT	09	Mortality	2015:12	TC	-3.75358	-280.5957
DE	10	Mortality	2012:11	LS	4.48328	88.7797
DC	11	Mortality	2002:01	AO	4.84057	115.7816
FL	12	Mortality	2000:03	AO	-4.13081	-1166.8559
FL	12	Mortality	2003:12	TC	4.35066	1155.6626
GA	13	Mortality	2003:12	TC	4.92570	638.9594

Table 1: **All Demographic Anomalies detected.** *(continued)*

state	STATEID	component	time	type	tstat	coefhat
GA	13	Mortality	2008:02	TC	3.53556	445.0709
GA	13	Mortality	2014:12	TC	4.61761	614.5234
HI	15	Fertility	2014:05	AO	4.96017	226.2289
HI	15	Mortality	2015:01	AO	3.63726	106.6684
ID	16	Fertility	2008:01	TC	3.77381	179.1547
IL	17	Fertility	2008:02	AO	3.50138	736.1977
IN	18	Mortality	2003:12	AO	4.86247	667.1098
IN	18	Mortality	2013:01	AO	5.89220	782.1618
IN	18	Mortality	2015:01	AO	5.08068	719.9057
IA	19	Mortality	2000:01	AO	4.22067	329.1083
IA	19	Mortality	2003:12	TC	4.18714	318.6955
IA	19	Mortality	2004:03	TC	-4.84868	-367.8987
IA	19	Mortality	2013:01	AO	4.24874	304.9240
KS	20	Fertility	2012:08	TC	4.86363	298.5049
KY	21	Mortality	2000:01	AO	4.39043	407.5624
KY	21	Mortality	2000:03	TC	-4.04460	-461.6888
KY	21	Mortality	2003:12	TC	5.56447	555.6897
LA	22	Fertility	2005:08	TC	-6.00213	-702.9202
LA	22	Fertility	2005:10	TC	-4.99691	-620.4862
ME	23	Mortality	2010:09	LS	5.05604	59.2778
ME	23	Mortality	2015:01	LS	7.15573	99.5227
MD	24	Mortality	2003:12	AO	3.95006	466.4458
MD	24	Mortality	2015:02	LS	4.20150	286.7422
MA	25	Fertility	2005:01	AO	-3.93933	-482.4588
MA	25	Mortality	2000:01	AO	5.26608	656.1995
MA	25	Mortality	2003:12	TC	4.59795	575.2061
MA	25	Mortality	2004:02	LS	-6.28503	-302.4497
MA	25	Mortality	2005:01	AO	4.78056	577.1766
MA	25	Mortality	2006:03	TC	3.69979	453.4105
MA	25	Mortality	2013:01	AO	6.16021	776.8701
MA	25	Mortality	2015:01	LS	8.68608	444.6978
MA	25	Mortality	2015:12	TC	-4.16043	-534.6610
MI	26	Mortality	2000:01	AO	5.10046	882.6971
MI	26	Mortality	2000:03	AO	-4.54718	-820.2133
MI	26	Mortality	2008:02	TC	4.34279	830.4293
MI	26	Mortality	2014:05	LS	5.54131	565.8971
MN	27	Mortality	2000:02	TC	-3.96853	-405.9615
MN	27	Mortality	2000:03	AO	-3.65185	-336.2720
MN	27	Mortality	2004:03	AO	-3.92523	-340.5204
MN	27	Mortality	2005:01	TC	4.56133	426.7506
MN	27	Mortality	2013:01	AO	4.73263	401.8875
MS	28	Fertility	2006:08	TC	4.29545	335.9598
MS	28	Mortality	2000:01	AO	3.60022	328.1801
MS	28	Mortality	2012:04	LS	11.13241	234.4396
MO	29	Mortality	2000:01	AO	6.70617	829.4624
MO	29	Mortality	2003:12	AO	5.72193	678.5393

Table 1: **All Demographic Anomalies detected.** *(continued)*

state	STATEID	component	time	type	tstat	coefhat
MO	29	Mortality	2005:03	AO	3.64903	433.0229
MO	29	Mortality	2008:02	TC	6.25163	738.5004
MO	29	Mortality	2012:06	LS	7.44381	285.5670
MO	29	Mortality	2016:04	LS	5.33701	276.1651
MT	30	Mortality	2005:02	TC	4.56533	131.9953
MT	30	Mortality	2016:12	TC	4.31327	136.2263
NE	31	Mortality	2003:12	AO	4.61562	219.5569
NE	31	Mortality	2004:03	TC	-3.94876	-185.1389
NV	32	Fertility	2004:05	AO	-3.75743	-284.9307
NH	33	Fertility	2008:02	LS	-5.68403	-75.5613
NH	33	Fertility	2010:02	LS	-5.38582	-73.6777
NH	33	Mortality	2010:04	LS	3.50759	49.1056
NH	33	Mortality	2014:11	LS	6.68259	94.5200
NJ	34	Mortality	2001:01	AO	-5.25394	-730.1265
NJ	34	Mortality	2001:09	AO	5.70305	733.1479
NJ	34	Mortality	2004:02	LS	-4.16336	-367.4317
NJ	34	Mortality	2005:01	TC	4.06743	592.5153
NJ	34	Mortality	2008:03	AO	3.70093	470.6901
NJ	34	Mortality	2011:01	TC	3.67652	512.7843
NJ	34	Mortality	2013:01	AO	6.68083	839.4441
NJ	34	Mortality	2015:01	TC	3.88918	572.1396
NM	35	Mortality	2008:02	AO	4.89007	232.9649
NY	36	Fertility	2009:08	LS	-3.81613	-675.9306
NY	36	Fertility	2013:02	LS	-3.73425	-685.2435
NY	36	Mortality	2000:01	AO	5.90125	1523.4348
NY	36	Mortality	2001:09	AO	6.39634	1627.5795
NY	36	Mortality	2004:02	LS	-5.59387	-868.8630
NY	36	Mortality	2005:01	AO	4.68886	1178.1655
NY	36	Mortality	2008:02	TC	4.49333	1313.1883
NY	36	Mortality	2013:01	AO	5.09972	1286.7644
NY	36	Mortality	2015:01	TC	3.90770	1168.8950
NC	37	Mortality	2000:01	AO	4.35640	638.6683
NC	37	Mortality	2000:03	TC	-4.13023	-727.6876
NC	37	Mortality	2003:12	AO	6.13518	792.9569
NC	37	Mortality	2006:03	AO	3.62848	465.8942
NC	37	Mortality	2012:04	LS	4.38439	418.6341
NC	37	Mortality	2015:01	LS	5.43649	525.6793
NC	37	Mortality	2017:03	AO	3.94487	574.9178
OH	39	Mortality	2000:01	AO	5.73968	1097.0115
OH	39	Mortality	2000:03	TC	-3.63496	-868.3683
OH	39	Mortality	2003:12	TC	6.90558	1528.0049
OH	39	Mortality	2005:03	AO	4.31174	812.7438
OH	39	Mortality	2008:02	TC	3.57092	773.9587
OH	39	Mortality	2013:01	AO	6.81518	1250.7794
OH	39	Mortality	2015:02	LS	-4.26921	-1157.3393
OK	40	Mortality	2000:01	AO	6.53041	564.1853

Table 1: **All Demographic Anomalies detected.** *(continued)*

state	STATEID	component	time	type	tstat	coefhat
OK	40	Mortality	2003:12	AO	6.12829	484.7089
OK	40	Mortality	2008:02	AO	5.72159	453.1802
OK	40	Mortality	2013:01	LS	5.88299	240.6262
OR	41	Fertility	2005:02	AO	-3.52100	-309.3496
OR	41	Mortality	2000:01	AO	4.45571	292.0508
OR	41	Mortality	2003:12	AO	3.64746	238.8606
OR	41	Mortality	2013:01	TC	4.11413	289.2114
OR	41	Mortality	2017:01	AO	6.98586	478.7094
PA	42	Fertility	2008:02	AO	3.78262	704.6169
PA	42	Mortality	2000:01	AO	5.21016	1287.6043
PA	42	Mortality	2003:12	TC	6.34267	1556.3416
PA	42	Mortality	2005:02	TC	3.75053	995.8809
PA	42	Mortality	2008:02	TC	4.35277	1092.9297
PA	42	Mortality	2013:01	AO	4.10701	939.2149
RI	44	Mortality	2000:01	AO	4.87114	207.5642
RI	44	Mortality	2005:01	TC	4.13923	155.9976
SC	45	Mortality	2000:03	TC	-4.18260	-395.5357
SC	45	Mortality	2003:12	TC	4.95743	428.4264
SC	45	Mortality	2005:02	TC	4.99622	449.9698
SC	45	Mortality	2008:02	TC	3.84015	346.2581
SC	45	Mortality	2011:12	TC	-4.33533	-386.8461
SC	45	Mortality	2016:02	AO	4.00147	360.7129
TN	47	Mortality	2000:01	AO	3.87715	515.2153
TN	47	Mortality	2003:12	AO	5.97624	757.8804
TN	47	Mortality	2008:02	AO	3.60115	449.4722
TX	48	Mortality	2000:01	AO	5.77140	1784.2714
TX	48	Mortality	2002:01	TC	5.28612	1481.2292
TX	48	Mortality	2005:02	TC	4.62638	1310.8300
TX	48	Mortality	2008:02	AO	5.04507	1441.3564
TX	48	Mortality	2013:01	AO	7.13433	1949.1723
TX	48	Mortality	2014:01	AO	4.47608	1229.5101
TX	48	Mortality	2015:01	AO	5.89546	1670.5812
UT	49	Mortality	2003:12	AO	4.18091	168.9354
VT	50	Fertility	2005:07	TC	-4.85815	-104.7377
VT	50	Mortality	2000:01	AO	4.19129	99.3281
VT	50	Mortality	2001:06	TC	3.61313	70.8473
VT	50	Mortality	2008:01	AO	-4.11211	-89.2642
VA	51	Mortality	2003:12	AO	7.05760	919.4141
VA	51	Mortality	2006:03	AO	3.96012	508.3156
WA	53	Mortality	2003:12	AO	4.53139	487.4918
WA	53	Mortality	2017:01	TC	6.00615	724.6422
WV	54	Mortality	2000:01	AO	3.92141	258.4560
WV	54	Mortality	2003:12	AO	5.38802	338.7782
WV	54	Mortality	2012:04	LS	5.31478	106.9373
WI	55	Fertility	2008:02	AO	3.65776	422.4784
WI	55	Mortality	2000:01	AO	3.86741	445.4985

Table 1: **All Demographic Anomalies detected.** *(continued)*

state	STATEID	component	time	type	tstat	coefhat
WI	55	Mortality	2005:02	AO	4.72419	536.8851
WI	55	Mortality	2008:02	TC	3.52011	379.2764
WI	55	Mortality	2012:07	LS	5.84049	345.4592
WI	55	Mortality	2013:01	AO	5.73090	638.0866
WY	56	Mortality	2002:01	AO	3.71787	89.7037

AL – Monthly Births **AZ – Monthly Births** **AL – Monthly Deaths** **AZ – Monthly Deaths**

CT – Monthly Births

DE – Monthly Deaths

CT – Monthly Deaths

DC – Monthly Deaths

FL – Monthly Deaths **HI – Monthly Births** **GA – Monthly Deaths** **HI – Monthly Deaths**

OR – Monthly Births

PA – Monthly Births

OR – Monthly Deaths

PA – Monthly Deaths

RI – Monthly Deaths

TN – Monthly Deaths

SC – Monthly Deaths

TX – Monthly Deaths