Video Games and Attitude Change: A Meta-analysis

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Author note

Data, R code documenting the analytic workflow, and analytic outputs (the supplementary material) are freely available at the Open Science Framework:
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Abstract

Despite extensive research on attitudes and a rapid growth of the video game market, there is currently no meta-analysis mapping the link between narrative video games and attitude change. Here, we present such meta-analysis. The findings suggest that narrative video games affect players' attitudes towards the topics depicted in games. This effect was present in studies focused on changes in both implicit ($g = 0.36$, $k = 18$) and explicit attitudes ($g = 0.24$, $k = 101$), with longer intervention duration and game mechanics such as stereotyping and meaningful feedback resulting in larger implicit attitude change. Regarding the robustness of the underlying evidence, half of the included studies were judged to be at high risk of bias. On the other hand, the impact of publication bias in this literature was found to be negligible. Altogether, this meta-analysis provides evidence that video games shape how we think about events they represent.

Keywords: attitudes, video games, meta-analysis, game mechanics, persuasion

Video Games and Attitude Change: A Meta-analysis

Tens of thousands of new video games are released every year (Statista 2021a; Statista 2021b; Grayson 2020) and just over one out of every three people on the planet is playing them (Newzoo, 2020). Video games necessarily address a number of topics we face in our daily lives - violence, stereotypes, history, ethnical conflicts and more. But are video games able to affect our attitudes towards the phenomena they deal with?

Despite growing empirical research on attitudes, there is currently no meta-analysis of video games' effect on attitudinal change. A previously conducted narrative review (Soekarjo & van Oostendorp, 2015) focused only on the effects of serious games and included just six studies. Five of those studies found a significant attitude change after playing the game; however, the review focused on articles examining pre-identified games. It excluded commercial games, and it reviewed games which are already 7 years old.

From the perspective of attitude change research, crucial potential lies in narrative video games. We view narrative video games as those with any type of narrative elements. These elements can be ones that add characteristics to a general object in the game. For example, features can be added to an object so that it visually and narratively represents a rainforest. As such, it has a certain impact on the game world and story and responds to players' actions. The way developers use and represent the rainforest's impact in the game is an example of how narrative elements can carry the meaning of game processes and actions that potentially affect our understanding and interpretation of displayed realities in video games. Therefore, narrative video games can deliver attitude-related information about particular topics and only these games can directly connect game actions' meanings with the realities we encounter in our everyday lives.

Due to the current, historically unprecedented rapid growth of the video game market and experimental video game development, there is a need for the first meta-analysis focused on the current state of empirical evidence regarding narrative video games and attitude change. By assessing 3,832 studies identified by our search operators in relevant databases, we deliver such a meta-analysis reacting to this research gap.

Using the Associative-propositional evaluation model (APE model; Gawronski & Bodenhausen, 2014; 2007; 2006), we investigate whether video games affect explicit and implicit attitude changes in the short term and over the long term. Considering further moderators of attitude change, we focus on the effects of intervention duration, persuasive game mechanics, the effect of action vs. non-action games, and comparison to various types of control groups.

Theoretical Background of Attitude Change

As defined by Vogel and Wanke (2016, pp. 2), attitude is a “summary evaluation of an object of thought. An attitude object can be anything a person discriminates or holds in mind”. The crucial characteristic of attitudes is their tendency to evaluate some attitude object with some degree of “favor or disfavor” (Eagly & Chaiken, 1993, p. 1). Attitudes are an essential factor when we process complex information (Sanbonmatsu & Fazio, 1990). They influence our information selection and the way we interpret obtained information (Pratkanis, 1989; Case & Given, 2016; Vogel & Wanke, 2016). Therefore, they play a key role in our interpretation of the world around us. The core mechanism to change someone’s attitudes is through processing information related to the attitude object (Crano & Prislin, 2008). However, when we confront information that is not aligned with our attitudes, it affects how we evaluate an information source’s credibility concerning that particular topic (van Strien et al., 2016). If the information provided is inconsistent with our attitudes, we consider the source less credible and vice versa

(from now on, we will call this phenomenon *credibility bias*). Perceived low credibility of the information source limits its persuasive potential.

According to the APE model, we distinguish explicit and implicit attitudes. These two forms of attitudes are guided by different, but often interplaying, processes (Gawronski & Bodenhausen, 2014; 2007; 2006). Implicit attitudes are derived from associative evaluations; that is, immediate affective responses to the object. Associative evaluations are based on the object's familiarity with other concepts in our memory, so the APE model assumes the existence of a mental structure containing these mental associations in the long-term memory. This mental structure can be changed by the co-occurrence of two concepts in one's environment resulting in either strengthening the associative link between these concepts or in creating a new associative link between them. For example, exposing participants to pictures of healthy, fit people eating a plant-based diet and to pictures of obese, unhealthy people eating an animal-based diet prior to an implicit attitude measurement towards different diets will result in temporarily more favorable attitudes towards a plant-based diet (see, e.g., Banaji & Greenwald, 2013). Seeing a plant-based diet regularly associated with positive perks in one's environment will result in the creation of a more permanent associative link between them in the long-term memory. Then, we can expect positive evaluation of the plant-based diet in implicit attitude measurements even without prior exposure to any stimuli: such as the pictures already mentioned. Implicit attitudes are assessed using response time measures. One's associative evaluations function independently in relation to what one consciously considers to be the truth.

On the other hand, explicit attitudes are based on propositional reasoning; that is, the logical conclusions derived from information related to the object in question. To change one's propositional reasoning about an object, one must be exposed to information which is not in line with one's current beliefs and/or with what one considers to be the truth. One cannot have two

contradictory propositional reasonings about the same topic as this would create cognitive dissonance and then need to be resolved. In such a case, one can either reject one of the propositions or seek additional information to resolve the cognitive dissonance created and the consistency of one's beliefs (Festinger, 1958). Explicit attitudes are very often assessed using self-reported questionnaires. The measurement of these changes is, however, limited to one's willingness or ability to share their explicit attitudes.

Narrative Video Games and Attitude Change

As has already been said, we consider narrative video games to be those with any type of narrative game elements. These elements provide particular meaning to processes and actions regarding depicted realities in the video games. For instance, a game can portray an electric car (compared to a car with an internal combustion engine) as a much more efficient and faster solution within the gameplay: either by how it is framed in the game story or by setting parameters for these cars. Narrative video games are often perceived as a source of entertainment, but they can also serve as a source of information. These games require players to seek and process information in the game narratives to proceed further in the game story or to fulfill game objectives. Therefore, narrative video games are information systems, in which players have the opportunity to react to the depicted information, experience the results of their actions, and respond to the changes caused in the game world (Kolek et al., 2021; Smethurst & Craps, 2014). As such, narrative video games can deliver messages about particular topics and connect game processes and actions to those in our everyday lives.

Persuasive Game Mechanics

Changes in implicit and explicit attitudes depend on different, and often interconnected, processes. Therefore, we expect that the different ways narrative video games represent information about particular topics, i.e., their persuasive game mechanics, will affect explicit and implicit attitudes differently. Persuasive game mechanics can affect attitudes towards the measured phenomena regardless of the game designers' intentions.

Stereotyping

Stereotyping is a persuasive game mechanic that constructs an implicit connection between some particular group of persons depicted in the narrative game elements and some stereotype trait assigned to that group (e.g., depiction of stereotypically represented Arabs as terrorists in an action game). This connection is usually not a carrier of essential importance for progression in the game, but it is still ubiquitous during gameplay. Also, the mechanic does not usually provide complex information about this connection. As such, we assume this mechanic will affect implicit rather than explicit attitudes.

Meaningful Feedback

As we define it, Meaningful feedback is a persuasive game mechanic that is directly related to progression in the narrative game. When players perform acts related to the measured phenomenon depicted in the narrative game, they are rewarded by a positive or negative outcome. This creates the connection between the act and its outcome value. Meaningful feedback can be implemented in various ways: including a simulation through an economical model depicting how something works; through representation of game rules and processes (Procedural Rhetoric; Bogost, 2010); through some form of a reward system linking some actions with positive rewards or to penalties; or through a combination thereof. For instance, if the

narrative game simulates a child picking cotton in a field as a tedious, hard and frustrating activity with a little reward, it creates a connection between child labor and these negative feelings in the gameplay. Meaningful feedback can affect either explicit or implicit attitudes as it can provide complex information about the outcome of some actions. It can also create frequent connections between measured phenomenon and some particular concepts that have positive or negative value for players.

Perspective-taking

Perspective-taking is a persuasive mechanic that provides players with a complex take on the measured phenomena from multiple points of view. Therefore, it introduces arguments about the topic: often from complementary or contradicting perspectives (e.g., providing players with interpretations of the Syrian conflict from the perspectives of all parties involved). Some arguments are in favor of, and some are against, players' initial explicit or implicit attitudes, but they all come from the one source. As such, they can possibly mitigate the credibility bias (see Section Theoretical Background of Attitude Change). We expect that perspective-taking will affect explicit rather than implicit attitudes.

Game Genres

For the sake of the present meta-analysis, we contrast just two categories of game genres: action games and non-action games. The reason for this split is as follows: Action games often put players in time pressure situations requiring their imminent reactions to proceed further in the game compared to the non-action games which do not. Also, success in action games relies more on hand-eye coordination and motor skills to control the game precisely in order to react to imminent stimuli in limited periods of time. We consider games that do not fit these characteristics to be non-action games. There is potentially a difference between narrative action

games and narrative non-action games in relation to attitudes. On the one hand, in narrative non-action games, players can have more time to perceive complex information about the depicted topics; and this with less distractions. On the other hand, narrative action games can create direct connections between the depicted topics in these games and some concepts of positive or negative value despite potential time pressure. The forming of these connections does not require transmission of complex information to happen. Therefore, narrative non-action games should be more suitable to affect explicit attitudes and narrative action games should be more suitable to affect implicit attitudes.

Hypotheses

H1: Narrative Video Games Induce a Change in Players' Explicit (H1a) and Implicit Attitudes (H1b)

Narrative video games can deliver messages about depicted topics and affect players' attitudes towards these topics within their persuasive game mechanics. They also allow players to interact with these depicted topics and, by doing so, let players challenge their own evaluations of the topics. There are several studies suggesting short-term and long-term effects of narrative video games on attitude evaluations (e.g., Kolek et al., 2021; Kampf, 2015; 2016a). Therefore, we assume that, on a general level, narrative video games will significantly affect explicit and implicit attitudes.

Next, we examine the following moderating effects.

Intervention Duration

H2: Duration of Intervention is Positively Related to the Magnitude of Explicit (H2a) and Implicit (H2b) Attitude Change. Longer duration of an intervention results in a longer period of time during which participants are exposed to information related to the measured

phenomenon within the game, i.e., how long they are exposed to the game's persuasive mechanics. On the most general level, to change someone's attitude means to provide this person with information in any form that will challenge their initial attitude. A study by Maier and Richter (2013) suggests that even attitude-inconsistent information will not be ignored (as assumed by the cognitive dissonance theory (Festinger, 1958)), rather it will affect our attitudes less than attitude-consistent information. Therefore, we assume that the longer intervention duration will affect players' attitudes (both implicit and explicit) more compared to the short intervention durations.

Type of Comparator

H3: The Magnitude of Explicit Attitude Change is Smaller in Studies Using Control Groups with Topic-related Activities than It Is in Studies Using Control Groups with Topic-unrelated Activities. We do expect that narrative video games have the potential to affect players' explicit and implicit attitudes as already elaborated. At the same time, there is no empirical evidence suggesting that activities unrelated to the measured phenomenon will affect explicit attitudes towards that phenomenon. Therefore, we expect that the magnitude of explicit attitude change in experimental groups in comparison to a corresponding control group will be smaller in studies using control groups with topic-related activities (e.g., books or lectures about the measured topic) than in studies using control groups with topic-unrelated activities (e.g., playing an unrelated game, watching an unrelated documentary).

H4: Narrative Video Games Change Player's Implicit Attitudes Only if Compared to Control Groups with Activities Unrelated to the Measured Topic (H4a), but Not if Compared to Control Groups Using Activities (H4b) Related to the Measured Phenomena. We do expect that the main mechanism causing implicit attitude change, i.e., co-occurrence of

two concepts in one's environment, can occur comparably in the game as in any related activity since it does not require any representation of complex information.

Game Mechanics and Their Effects on Attitudes

Implicit and explicit attitude changes depend on different and often interconnected processes. Therefore, we also expect that the different persuasive game mechanics affect explicit and implicit attitudes differently (see Table 1), as detailed in the following text.

H5: Implicit Attitudes Are More Affected by Games Using Stereotyping and Meaningful Feedback as Persuasive Game Mechanics than by Other Game Mechanics.

According to the APE model, mental structures responsible for our implicit attitudes can be changed by the frequent co-occurrence of two (or more) concepts in one's environment. We assume that this frequent occurrence of two concepts is more likely to happen when the game uses Stereotyping and Meaningful feedback as persuasive game mechanics. This is because they both can frequently link two (or more) concepts together within the game. The first one links the measured concept with some stereotypical characteristic; the second one links the measured concept with a positive or negative value assessment of players' actions within the game. Perspective-taking as a persuasive mechanic does not create such ubiquitous and clear connections between the two concepts within the gameplay.

H6: Explicit Attitudes Are More Affected by Games Using Perspective-taking and Meaningful Feedback as Persuasive Mechanics than by Other Game Mechanics. Changing explicit attitudes relies on the acquisition of new information that is potentially challenging one's current beliefs. On the empirical level, several studies suggest that Perspective-taking is able to affect explicit attitudes (e.g., Kolek et al., 2021; Todd & Gallinsky, 2014; Kampf, 2015; 2016a). Within the gameplay, Perspective-taking provides players with a complex, multi-perspective take on the measured phenomena. Furthermore, it should be able to mitigate credibility bias (see

Section Theoretical Background of Attitude Change). Therefore, we assume that Perspective-taking can affect players' explicit attitudes.

Meaningful feedback as a persuasive mechanic provides players particular freedom to do various, in-game actions related to the measured phenomenon. It assigns the outcomes of these actions a negative or positive value. The reasoning behind this assignment can be comprehensively elaborated within the game narrative. Therefore, we assume that the Meaningful feedback mechanic can affect players' explicit attitudes.

Game Genres

H7: Narrative Action Games Have a Larger Effect on Implicit Attitudes than on Explicit Attitudes. Narrative action games are more suitable for frequently demonstrating a particular link between two concepts (thus the mechanism affects implicit attitudes) than for providing topic-related complex information (thus the mechanism affects explicit attitudes). Therefore, we assume that implicit attitudes will be more affected by narrative action games than will be explicit attitudes.

H8: Narrative Non-Action Games Have a Larger Effect on Explicit Attitudes than Do Action Games. Attitude challenging information is required to affect someone's beliefs and thus their explicit attitudes. Narrative, non-action video games represent a format with a higher potential to demonstrate complex information to players without time pressure limiting the exploration of their narratives. Therefore, we expect that explicit attitudes will be more affected by narrative, non-action video games than by action games.

Exploratory Analyses

We also have two exploratory goals. First, we examine whether the magnitude of explicit attitude change remains the same over time between the immediate posttest and the delayed

posttest or whether it changes. Predictions from theories as regards long-term attitude change are mixed. On the one hand, as described by cognitive dissonance theory (Festinger, 1958), newly acquired information, which is not in line with one's explicit attitudes, will result in the creation of cognitive dissonance and the subsequent motivation to solve it. This can happen either by acquisition of new information, which will resolve the dissonance and change beliefs, or by rejecting the new information. Therefore, this theory assumes that one's explicit attitudinal changes will remain the same over the long term, or that they will, to some extent, shift back to their original values. On the other hand, the "sleeper effect" theory suggests that the persuasive effect can increase over time (Priester et al., 1999) when the message is thoughtfully elaborated. Within this meta-analysis, we intend to explore the direction of the long-term trend of explicit attitude change.

Second, we examine whether there is any significant effect of age or education on attitude change. Here, we do not have any particular prediction.

Method

Inclusion Criteria

The process of study selection is outlined in the Prisma Flowchart (Page et al., 2021; see Figure 1). We have included in our meta-analysis only studies that met the inclusion criteria in the following four domains:

Intervention

We have included only narrative video games as explained in Section Narrative Video Games and Attitude Change. We have also excluded studies focused only on particular game elements not related to the game narrative (e.g., game based-learning elements). However, studies dealing with the effects of game elements possessing any potentially significant meaning for a

game narrative, e.g., visual design or a dress-code for game characters, were considered as relevant for our analysis as they can be a part of persuasive game mechanics (see Section Persuasive Game Mechanics)

Outcome

We have included all studies dealing with attitudes, as defined in Section Theoretical Background of Attitude Change, and the changing thereof. This included all effects, impacts, changes or evolutions of attitudes in relation to the experimental intervention; i.e., playing the game. For the purpose of this meta-analysis, attitude change is considered the dependent variable. Furthermore, attitudes examined in the study had to be related to particular substantive topics represented in the game narrative. Therefore, we excluded all studies focusing on general attitudes towards a) games or playing games (e.g., Zhu, Lin & Hsu, 2012; Garneli, Giannakos & Chorianopoulos, 2017) or b) towards any other general activity related to the actual playing of the game; e.g., attitude towards competitiveness (Williams & Williams, 2011) or learning attitudes (e.g., Lin, Tutwiler & Chang, 2011).

Study Design

Studies in our research sample had to collect quantitative data and allocate subjects to the interventions and some sort of control group(s). At least one experimental group in the study had to experience intervention through a narrative video game. Beyond this general condition, the intervention by the narrative video game should have been the only intended element affecting a player's attitudes towards the topic. Based on this reasoning, we excluded studies with video game interventions that were preceded or accompanied, for example, by a seminar, workshop or collective debate about the topic (Strawhacker et al., 2018; Hornung et al., 2000) and also studies with external elements within the study design, e.g., driving a car in a video game while

telephoning in a real life (e.g., Downs, 2014) that would purposefully affect players. Along the same line of reasoning, we also excluded so-called “exergames”, e.g., games combining playing with physical activity as a form of exercise. Second, we also excluded studies examining the effectivity of “advergames”; that is, games promoting a product or a brand. Those games’ persuasiveness is rooted in their interest in favoring a particular commercial product or brand, which is a qualitatively different factor compared to other games in our study.

Data Availability

We also included three practical limitations. First, the paper had to be in English. Second, it had to contain the relevant data about the examined groups or be available upon request (all authors were contacted at least twice). Third, the paper’s full text had to be accessible (paywalled papers included).

Search Strategy

To optimize for the best tradeoff between the search’s recall and precision, we used an adaptation of the relative recall technique (Sampson et al., 2006). The strategy was to carry out a pilot search with maximum sensitivity, arrive at a legacy set of reference studies, and iteratively adjust the search string to return all or nearly all of the identified legacy set in each database while optimizing for specificity.

First, in September 2018, we pilot-searched Scopus, Web of Science Core Collection and Google Scholar databases using the search string “attitude* AND game*”. This broad pilot search was intentionally aimed at maximizing sensitivity, while sacrificing search specificity. We screened the first 1,000 most relevant studies in Google Scholar; the first 1,500 most relevant studies in Web of Science; and the first 500 most relevant studies in Scopus. We also examined references in the identified records to find information about other relevant studies. We ended up

identifying 26 possibly eligible studies. Second, we iteratively adjusted the pilot search string to make it more specific in order to recall all, or nearly all, of the 26 identified legacy studies present in each database. We did this while limiting the maximum number of hits per database to 1,500. Databases used to search the literature are listed in Table S1 in the Appendix A. The final search took place on August 18, 2020. Search capabilities differ slightly across databases, so we had to use a distinct translation of our search string for each database. Here is a search operator used for the Scopus database:

TITLE-ABS-KEY ((attitude* OR stereotype*) AND (change OR effect OR significant*¹ OR impact) AND (game*) AND (experiment* OR empirical* OR intervention)).

Selection of Studies

Using various translations of the above-given search string across the 8 databases (Table S1 in the Appendix A), we have identified 3,832 studies. For an additional eligibility check, we have included 34 possibly eligible studies from the pilot search not recalled during the systematic database check, citation searching, and other resources; i.e., one study in review (See Prisma flowchart in Figure 1).

Data Extraction

Coding quality was checked by utilizing a second coder who coded a random 30% of all studies included. Coding disagreements were discussed and, if needed, resolved by consulting one of the other authors. The aim was not just to catch coding errors, but also to look for potential problems in the coding scheme. One particular variable was redefined after this process – Persuasive mechanics. Originally, it consisted of 6 values – Stereotyping, Perspective-taking,

¹ In the pilot testing of the search string, the keyword “significant” helped to increase the recall of quantitative studies. This was the initial reason for including this keyword, but we also carried out the search with this keyword left out. This led to the omission of two studies, but there was no noteworthy change in the results ($\Delta g = 0.0051$). Details about excluding the keyword “significant” can be found in the Appendix B).

Economical model, Procedural rhetorics, Reward system and Others. However, the analysis revealed that the categories Economical model, Procedural rhetorics and Reward system overlapped and difference between them was not reliably recognizable. Therefore, we have decided to unite them into one category Meaningful feedback: due to the many similarities between them. Inter-rater reliability for metric variables ranged from Cohen's $\kappa = .67$ (for Intervention Duration) to $\kappa = 1$, with a mean Cohen's κ at .92. For categorical (mostly binary) variables, we computed the percentage agreement, which ranged from 76% (whether the effect was focal) to 100%, with a mean percentage agreement of 94%. Complete report of inter-rater agreement calculations can be found in the Appendix C (online only).

In case of missing data needed for the computation of effect sizes, we have contacted article authors at least twice by email, leaving at least a 14-day period between attempts.

Moderators

Type of Attitudes

We have distinguished between a) implicit attitudes, i.e., those assessed using response time measures (e.g., Implicit Association Tests); b) explicit attitudes, i.e., those assessed using self-reported questionnaires involving Semantic differentials or Likert Scales.

Intervention Duration

We have collected data about the means of experimental intervention durations. In cases where authors stated this value as an interval, we have used its mean value.

Persuasive Mechanics

We have divided the studies into four categories: a) Perspective-taking; b) Meaningful feedback; c) Stereotyping (see Section Persuasive Game Mechanics); and d) a Non-defined category for studies with unclear or multiple persuasive mechanics.

Posttest Delay

We have recorded data about the number of days between the intervention and the posttest collection of data from participants. In cases where authors stated this value as an interval, we have used its midpoint.

Game Genres

The games used in experimental interventions were coded as being divided into the following two categories: a) action games; b) non-action games (see Section Persuasive Game Mechanics for the theoretical background).

Control Groups

We coded control group types as follows: a) Activity unrelated, which involves a lecture/presentation/reading on a topic unrelated to the one in the experimental group; b) Game with a different mechanic, which is a game featuring an unrelated topic and using different mechanics than the game in the experimental group; c) Game with a similar mechanic, which is a game on an unrelated topic but which uses similar mechanics to the game in the experimental group; still, some of their aspects differ like avatars, mission, etc.; d) Activity- related, which involves a lecture/presentation/reading on a topic related to the one in the experimental group; e) No activity apart from study measurements and f) Combination of various activities, such as games, videos or reading.

Age/Education

We have collected data on participants' mean ages and also data on education level. We have distinguished the following categories: a) elementary school students; b) secondary school students; c) university students; d) other, which included everyone else outside the first three categories.

Other Moderators

Moderators that were coded but not used for any hypothesis are described in Appendix D (online only).

Effect Size Computation

We used primarily group posttest means, *SDs* (or *SEs*) and *Ns* to compute Hedges' *g*, a standardized mean difference effect type corrected for small sample bias (Hedges & Olkin, 1985). In case group descriptives were not available, we converted the effect sizes from reported test statistics or other types of effect sizes. The computation and conversion of all effect sizes were carried out in code, using formulas laid out in Borenstein et al. (2009). To counteract a possibly biasing effect from undisclosed subject exclusions, we checked whether the sum of group *Ns* approximately matched the total sample size ($N \pm 2$). If it did, we used the respective group *Ns*. If it did not, we tried to compute group *Ns* based on the reported degrees of freedom, assuming a balanced design. If only the total sample size was reported, we also assumed a balanced design. We excluded effects for which the essential data was not reported and could not be recovered from the authors.

Analysis

Effect sets including more than 10 effect sizes were considered informative and were synthesized employing multilevel random-effects models with Satterthwaite's small-sample adjustment. We included all theoretically relevant effects for each study. To account for dependencies among the effects, we employed robust variance estimation (RVE) with the CHE working model (Correlated and hierarchical effects; Pustejovsky & Tipton, 2020). These models account for both types of dependencies among the effects simultaneously – nesting of effects within studies and clustering due to estimation of effects based on the same participants. As data on the sampling correlations among the effects is frequently unavailable, a constant sampling

correlation of .5 was assumed. As a sensitivity analysis, we relaxed this assumption by varying the sampling correlation from 0 to .6 in increments of .2. To test for equality of effect sizes across the levels of the moderators studied, we used the robust HTZ-type Wald test (Pustejovsky & Tipton, 2020).

Apart from the effect size estimates, we examined the absolute and relative heterogeneity using τ and I^2 , respectively. To estimate the range of true effects to be expected in similar future studies, we calculated 95% prediction intervals.

Prior to our analyses, we carried out an in-depth diagnosis of the random-effects meta-analytic model. Specifically, we screened for influential outliers using the Baujat plot and influence diagnostics indices. Outliers exerting an excessive influence on the meta-analytic model (if any) were only excluded in a sensitivity analysis.

In a sensitivity analysis, we also checked whether excluding studies with a high overall risk of bias (utilizing algorithmic-based judgment) and effects based on mathematically inconsistent means or SDs did have a meaningful influence on the meta-analytic inferences.

All models were fitted using restricted maximum-likelihood estimation using R packages metafor, version 2.5 (Viechtbauer, 2010) and clubSandwich, version 0.4.2. (Pustejovsky, 2020). The data analysis was carried out in R also using the following packages: esc (Lüdtke, 2017), tidyverse (Wickham et al., 2019), lme4 (Bates, Maechler, Bolker, Walker, 2015), dmetar (Harrer et al., 2019), and psych (Revelle, 2018).

Adjustment for Publication Bias

As null or negative results are less likely to get published, available studies represent a biased sample of the conducted (and all conceivable) studies. Under the influence of publication bias, the meta-analytic effect size estimates tend to be inflated to an unknown and possibly

substantial degree and have an excessive false-positive rate (Carter et al., 2019; Hong & Reed, 2020; Ioannidis, 2008).

Although bias adjustment methods assume a more realistic selection process, they may fail to recover the “true” magnitude of the studied effects under a number of realistic conditions. The estimates should thus rather be seen as approximations (see Ropovik et al., 2021). If the adjusted estimates from selection models markedly diverged from the crude meta-analytic estimates, then we primarily used bias-corrected estimates to guide our substantive inferences.

Selection Models

As the primary bias-adjustment approach, we applied a permutation-based implementation of the step-function selection model (see McShane et al., 2016). Selection models are a statistically principled, highly flexible family of models that directly map the functional form of the biasing selection process. In short, a 3-parameter selection model includes the following parameters: population effect size, heterogeneity, and the likelihood that a non-significant vs. significant result gets published. The model then uses maximum likelihood to estimate the three parameter values under which the observed data are most likely (McShane et al., 2016). By default, we applied the 4-parameter selection model (it also estimates the probability of the effect being in the opposite direction). If there was too little data (at least one of the p -value intervals contained less than 4 focal p -values), the estimation procedure automatically reverted to the 3-parameter selection model. All selection models (including the one-parameter selection models p -uniform* and p -curve) subset only the results that were deemed to be the study’s focal effects (reported in the abstract).

As selection models suited for the analysis of multi-level data are yet to be developed, the dependencies among the effects were handled using a permutation-based approach. We randomly

drew only a single focal outcome from each study, estimated the model repeatedly in 5,000 iterations, and averaged over this set of iterations by taking the model with the median estimate. This procedure sidesteps the use of arbitrary and potentially biasing decision rules for choosing independent effects.

To examine the variability in adjusted effect size estimates under different assumptions about the selection process, we also computed a series of three Vevea and Woods (2005) step function models with a priori defined selection weights. We used a fine-grained 10-step function to model different levels of severity of bias: moderate, severe, and extreme.

Exploratory Bias-adjustment Methods

For exploratory purposes, we also supplemented the primary selection modeling approach with the following secondary methods. Namely, we used the multi-level, RVE-based implementation of the PET-PEESE method (Stanley & Doucouliagos, 2014), Weighted Average of the Adequately Powered studies (WAAP-WLS estimator; Stanley, Doucouliagos, & Ioannidis, 2017) and p -uniform* (van Aert & van Assen, 2021). The details about the implementation of these methods with the results of the latter two methods are provided in the Appendix E).

Quality of Evidence Assessment

To appraise the quality and integrity of the evidence, we have carried out the following procedures.

Risk of Bias

First, we have assessed the risk of bias using the Revised Cochrane risk of bias tool for randomized trials (RoB 2; Sterne et al., 2019). The risk of bias was assessed in five domains: namely, bias arising from the randomization process; bias due to deviations from intended interventions; bias due to missing outcome data; bias in measurement of the outcome; and bias in

selection of the reported result. The judgments about bias in these domains were made using an algorithmic approach based on signaling items. When justified, the assessor could override the suggested risk of bias judgments, but this could be done only conservatively, i.e., in the direction of downgrading the judgment.

Numerical Inconsistencies in Reported Means and SDs

Second, using GRIM (Brown & Heathers, 2017) and GRIMMER (Anaya, 2016) tests, we tried to identify effects based on means or standard deviations that are mathematically inconsistent with the reported sample sizes. Checking for such inconsistencies is possible if the outcome was a discrete variable (e.g., Likert-type individual items or scales). In that case, means and *SDs* follow a fixed granular pattern for each combination of *N* and the number of items (Anaya, 2016; Brown & Heathers, 2017).

Numerical Inconsistencies in Reported *p*-values

Next, we screened all included studies for inconsistencies in reported *p*-values. This machine-based screening was carried out using the statcheck package (Epskamp & Nuijten, 2018). The method works as follows: (1) pdf files are converted to plain text, (2) which gets scanned for statistical results reported in APA style, (3) test statistics and degrees of freedom are extracted to recompute the *p*-value, (4) which is compared to the reported *p*-value. Having extracted that data, we computed in which proportion of cases the *p*-values were inconsistent with the reported test statistics and how many of those cases led to an inferential decision error.

Assessment of Evidential Value

Using the *p*-curve method, we tested whether selective reporting can be ruled out as the sole explanation of the observed findings (Simonsohn et al., 2014). If there is evidential value in the given literature, a right-skewed distribution of *p*-values can be observed regardless of power.

It follows that a set of direct replications is expected to yield a non-zero effect. On the other hand, a left-skewed distribution of p -curves may indicate a substantial prevalence of questionable research practices in the literature.

In the present meta-analysis, p -values were recomputed from the reported descriptive statistics. The dependencies between the p -values were handled using a permutation-based procedure, repeatedly drawing only a single focal effect from each study (with 200 iterations), estimating the p -curve, and averaging over the set by selecting the model with the median z -score for the right-skew of full p -distribution.

Median Statistical Power in the Literature

Lastly, we also computed the average statistical power to detect various smallest effect sizes of interest (0.20, 0.50, and 0.70). In the supplementary materials, we also report median power to detect the bias-corrected estimates.

Results

Sixty-seven studies from 40 papers matched the inclusion criteria. Out of those, 58 independent-sample studies (reported in 35 papers) provided sufficient information to recompute 119 effect sizes. In total, the included set of effects summarized data from 14,272 unique participants, with a median N across the included effects of 127. The vast majority of the included effects (96%) originated from randomized studies (see Appendix F for the full analytic outputs and Appendix G for the list of studies and the effect sizes; both online only).

Prior to the analyses, we screened the full meta-analytic set of 119 effects for outliers. Based on the examination of the Baujat plot and influence diagnostics indices, none of the included effect sizes exerted undue influence on the meta-analytic model.

In what follows, we first carry out a comprehensive synthesis of the entire literature that provides evidence on video games' overall effect on attitude change. Second, we address the substantive questions posed by this review. Third, we look at the methodological moderators to identify design-related and meta-study factors that may affect the size of the detected effects in this literature. Fourth, we conduct a detailed appraisal of the quality of empirical evidence; check for the presence of reporting inconsistencies; and indications of *p*-hacking at the level of the literature. Lastly, we carry out several sensitivity analyses to examine the robustness of our results to arbitrary methodological decisions.

Narrative Video Games' Overall Effect on Attitude Change

The set of effects reported in the literature that concerned narrative video games' effect on attitude change exhibited a high degree of heterogeneity, $Q(118) = 483, p < .001$. The standard deviation of true effects was $\tau = .40$, while $I^2 = 93\%$ of the total variance across the observed effect estimates was of a systematic nature (86% due to between- and 7% due to within-cluster heterogeneity). Accordingly, the 95% prediction interval was wide; with the true effect in a newly published study being expected to fall between -0.55 and 1.05.

The Random-effects RVE-based model estimated a mean effect size of $g = 0.25$, 95% CI (0.14, 0.37), $p < .001$, which is a small effect. Publication bias-adjusted effect estimated by the permutation-based 3-parameter selection model (3PSM) was likewise significant and of similar magnitude as the unadjusted estimate, $g = 0.33$, 95% CI (0.14, 0.53), $p = .001$. To further examine the variability in bias-adjusted estimates under different assumptions about the selection process, we also computed a series of Vevea & Woods (2005) step function models with a priori defined selection weights. The effect size estimates for the assumed moderate, severe, and extreme selection of results for publication were 0.15, 0.04, and -0.07. The markedly higher estimate for the maximum-likelihood-based 3PSM thus indicates that the selection by

significance in the given literature is less severe than the selection process assumed by the “moderate” model. In fact, only 14% of all effects included from primary studies were significant and Kendall’s correlation between the effect sizes and their standard errors was only $r = .04$ – suggesting only slight asymmetry in the chances of non-significant and significant effects to be published (as can also be seen from the funnel plot in Figure 2).

As a last, exploratory approach to bias correction, we also applied the multi-level RVE-based implementation of the PET-PEESE method. Assuming a hypothetical study with an infinitely large sample size, the method did not indicate the presence of an effect and returned a bias-adjusted effect size estimate that was effectively zero, $g_{PET} = -0.03$, 95% CI $(-0.34, 0.27)$, $p = .84$. The estimate was, however, rather weakly informative (judging by a relatively large CI width) and overlapped with the 3PSM estimate and thus was not significantly different. Although this exploratory result adds a layer of uncertainty by pointing to the rather suboptimal amount of information in the data, the primary analyses (naïve meta-analytic model and 3PSM) indicate the presence of a small, but robust, general effect of narrative video games on attitude change.

H1: The Effect of Narrative Video Games on Change in Explicit and Implicit Attitudes

A far larger proportion of the included studies examined the effect of video games on explicit attitude change ($k = 101$, 10% significant) than on implicit attitude change ($k = 18$, 39% significant). We did not detect substantial heterogeneity in the implicit attitude effects (neither absolute, nor relative), while the heterogeneity of explicit attitude effects was substantial (see Table 2). Both sets of effects yielded small-to-medium-sized average effects. Likewise, our primary bias-adjustment method (3PSM) indicated that, even after accounting for publication bias, the effect size estimate did not diverge from the unadjusted estimates. The pattern of estimates for the series of Veeva & Woods selection models was also lower than the naïve

estimates similar to the overall effect results. This suggests low severity of publication bias in both subsets. As the secondary, exploratory bias-adjustment method, PET-PEESE, detected a significant effect only for implicit attitude change (but not for explicit attitude change) of a practically identical magnitude as for the unadjusted effect. Overall, the literature we studied provides empirical evidence for a modest efficacy of video games for change of both explicit as well as implicit attitudes (based on the secondary PET-PEESE analysis, the effect on implicit attitudes seems more empirically robust though). H1 was thus supported. For more detailed results and plots, please see the full analytic output in the supplementary materials.

To compare the mean effects bound to explicit vs. implicit attitude change, we tested a meta-regression model controlling for several design-related factors that may have been prognostic with respect to the effect sizes (i.e., might vary between these sets of effects). We adjusted the comparison for overall risk of bias, published status, mean age of participants, and whether the intervention was administered in a lab. We did not find a difference between the effects related to explicit vs implicit attitude change, Wald's-type test $F(1, 3.13) = 0.45, p = .55$. Nor was there an effect with the covariates left out.

Substantive Moderators

H2: The Relation between the Duration of Intervention and Attitude Change

We did not detect a relationship between the duration of intervention and attitude change in the overall set of effects: meaning lack of support for H2 (Table 3). When broken down, however, effect sizes related to explicit attitudes (H2a) did not prove to be associated with the duration of the intervention ($p = .19$), while effects related to implicit attitudes (H2b) did ($p = .003$). When we subset just the effects based on delayed posttests, the magnitude of explicit

attitude change also did not change significantly over time (from immediate posttest to the delayed posttest).

H3 & H4: Characteristics of the Comparator Group and Attitude Change

We also first assumed that the magnitude of explicit attitude change (H3) is smaller in studies using control groups with topic-related activities than in studies using control groups with topic-unrelated activities. Although the difference between the respective subgroups was in the expected direction, $g = 0.29$ for subgroup using unrelated activities and $g = 0.13$ for related activities subgroup, it was not significant (Table 3). This indicates a lack of support for H3.

At the same time, we expected that video games would change player's implicit attitudes (H4) only if compared to control groups with unrelated activities towards the measured topic, but not if compared to control groups using related activities towards the measured phenomena. This prediction (as stated in H4) could not be tested as none of the implicit attitude effects were based on a design with controls doing related activities.

H5 & H6: Persuasive Mechanics and Attitude Change

Next, we tested whether implicit attitudes are affected more by games using Stereotyping and Meaningful feedback as persuasive mechanics than by Perspective-taking (H5). Games using Stereotyping and Meaningful feedback were associated with significantly larger effect sizes, $g = 0.37$ compared to games employing Perspective-taking, $g = 0.17$. The difference was statistically significant, thus corroborating H5.

On the other hand, we assumed that explicit attitudes are affected more by games using Perspective-taking and Meaningful feedback as persuasive mechanics than by other game mechanics, i.e., Stereotyping [H6]. While the pattern of mean effect sizes for these subgroups

was in the opposite direction, the difference was not significant. H6 was thus not supported by our data (see Table 3).

H7 & H8: Action Games' Effect on Attitude Change

Concerning the type of game, we expected that action games have a larger effect on implicit attitudes than on explicit attitudes (H7). Although the pattern of meta-analytic estimates was in the expected direction, $g = 0.32$ for implicit and $g = 0.23$ for explicit, the difference was not statistically significant.

From a different perspective, we also hypothesized that non-action games have a larger effect on explicit attitudes than do action games [H8]. Here, both subgroups were quite similar in terms of average effect size, with $g = 0.26$ for non-action games and $g = 0.21$ for action games. This difference was therefore not significant. Current data provided evidence neither for H7, nor for H8 (see Table 3).

Basic Characteristics of the Sample and Attitude Change

As an exploratory analysis, we also looked at the relationship between attitude change and (1) mean age of the participants and (2) the sample's gender composition (see Table 3). First, we found some feeble evidence of a negative relationship between the magnitude of attitude change and age. The impact of video games very slightly diminished with increasing age. On the other hand, the effect of video games seemed invariant with respect to the sample's gender composition (percentage of females).

Methodological and Meta-Study Moderators

In brief, we also assessed the moderating role of several design-related and meta-scientific factors.

First, we did not detect a difference between effects from non-laboratory studies ($k = 23$, $g = 0.15$) compared to effects coming from in-lab studies ($k = 88$, $g = 0.28$), $F(1, 5.68) = 1.05$, $p = .35$.

Second, studies restricting the sampling frame using demographic factors ($k = 97$) that may play a role in attitude change outcomes (thus decreasing the sampling variability) found larger effect sizes, $g = 0.31$, than studies not applying a restrictive sampling scheme ($k = 20$), $g = -0.04$, $F(1, 9.64) = 13.50$, $p = .005$. Since the majority of studies used some kind of range restriction, effect sizes found in this meta-analysis will likely prove to be smaller in more general samples. That said, an F -test of the equality of variances could not reject the hypothesis that the population variances for restricted and unrestricted samples were identical, $F(19, 97) = 0.90$, $p = .58$. Thus, the effect of range restriction may well be negligible.

Third, although effects produced by commercial games ($k = 51$, $g = 0.35$) were larger in our sample of effects, compared to non-commercial games ($k = 65$, $g = 0.18$), the difference (or the precision of the estimates) was not large enough to be significant, $F(1, 51.7) = 1.89$, $p = .18$.

Lastly, using covariate-adjusted models, we examined whether (a) the precision of the study designs has been improving over the years, (b) whether more informative (lower SE) studies tend to attract more citations, or (c) the same studies tend to get published in higher-impact journals, (d) whether studies reporting larger effect sizes tend to get more attraction, and lastly (e) whether there is a decline effect where studies showing more extreme (possibly opposite) results appear early in the research line rather than later as data accumulates (Ioannidis & Trikalinos, 2005). We found empirical support only for (c) the positive relationship between the study's precision and the journal's impact. Details and results of these analyses can be found in the supplementary materials.

Assessment of the Quality of Evidence Underlying the Overall Effect

Most concerns regarding the risk of bias in the included set of studies were due to bias arising from the randomization process (49% of studies being at low risk of bias) and due to bias in the selection of the reported results (only 20% being at low risk). Overall, only 29% of studies were rated as low risk, 15% raised some concerns, while 56% were at high risk of bias (i.e., being at high risk of bias in at least one domain or raising some concerns in multiple domains). See Figure 3 for more detailed results on each domain.

Second, 71% of the included effects targeted outcomes measured on a discrete scale. To examine the presence of reporting inconsistencies in the literature, we checked whether means and standard deviations underlying these effects were mathematically consistent with the reported sample sizes. Here, we found that 22% of these effects were flagged as being based on at least one mathematically impossible mean (18%) or *SD* (4%) – that is, mean or *SD* that cannot be arrived at given the reported *N*.

Third, we carried out machine-based, full-text screening (Epskamp & Nuijten, 2018) to extract all test statistics and the associated *p*-values reported in all 40 papers included (not just the ones providing sufficient effect size data). These test statistics were properly reported in APA style in 48% of the papers. Out of 469 extracted results, 10.5% were flagged as errors. In 8% of those flagged as errors, the error led to the opposite conclusion regarding the presence of the effect. Overall, reporting errors related to inconsistency between the reported test statistics and the *p*-value were distributed relatively evenly across the literature: with at least one such error being present in 58% of all papers included.

Fourth, we examined the indications of the presence of *p*-hacking using a permutation-based *p*-curve analysis. The distribution of *p*-values was right-skewed, indicating the presence of evidential value, $z_{\text{half}} = -2.74$, $p_{\text{half}} = .003$. As the *p*-curve only includes independent significant effects, the median model was based on just 7 effect sizes. That said, we did not detect any

pattern consistent with a large prevalence of *p*-hacking in this literature. This also held for both subgroups, implicit attitudes as well as explicit attitudes subsets, where both *p*-value distributions were associated with $p_{half} < .005$. The median *p*-curve for the overall effect can be seen in Figure 4.

Lastly, the overall literature was adequately powered (.98) to detect effect sizes of medium magnitude ($d = 0.50$) on average. On the other hand, the median statistical power to detect a small hypothetical effect size ($d = 0.20$) was relatively low; at only .36, on average.

Discussion

General Effect of Narrative Video Games on Attitudes

On the most general level, available evidence indicates that narrative video games do affect attitudes towards the depicted topics: both explicit and implicit. Video games are a popular phenomenon in our societies. The fact that they seem to have a marked impact on attitudes is crucial for debates about the way they represent our world. For instance, narrative video games' representations of marginalized groups, history or gender stereotypes are, according to these findings, transcending the medium itself and affect our daily lives and our interpretation of the world.

Beyond this general finding, we have also examined multiple moderators of this global effect: like intervention duration, different game design models embedded in video games, and other methodological aspects of research designs utilized in this field.

Intervention Duration

Evidence related to Hypothesis 2 was rather mixed. The intervention duration did not prove to have the presumed effect on explicit attitudes (H2a). Nevertheless, the longer intervention duration resulted in larger effect sizes on implicit attitudes (H2b) as presumed.

This means that the longer one plays the narrative video game, the larger the impact it has on their implicit attitudes but not necessarily on their explicit attitudes. These findings support the idea theoretically posited by the APE model: that implicit attitude change is caused by the frequent co-occurrences of the measured concept with another concept of positive or negative value. As exposure time seems to be related to the magnitude of the effect on associations about conveyed topics, this implies that narrative video games with longer durations may have a more significant effect on players' attitudes. Furthermore, it is important to mention that intervention durations in the analyzed studies are relatively short compared to dozens of hours of gameplay offered by most of the popular narrative video games on the market.

Possibly, as regards the lack of support for H2a, the effect on explicit attitude change may be tied to the role of participants' relationships to the measured topic. Unlike implicit attitudes, explicit attitudes are based on multiple, consistent, logical conclusions which reflect what one considers to be truth. As such, the participants' relationships to the topic is potentially of similar or greater importance than the intervention duration. This has been suggested by a few studies. For example, the results of studies by Kampf (2016b) and Alhabash & Wise (2015) indicate that belonging to a particular national group is crucial for the effect of narrative video games on attitude change towards the Israeli-Palestinian conflict or the parties involved.

The finding that the length of the intervention might not be essential for changing explicit attitudes is also supported in the study by Pena et al. (2018). They collected data in the middle of the experiment and at its end; participants' explicit attitudes towards the topic did not change between those two data collection points even though the duration of the intervention was doubled. However, more data is needed on that issue.

Persuasive Mechanics

Our results indicate that particular persuasive game mechanics (Stereotyping and Meaningful feedback) have a larger effect on implicit attitudes than others (H5). On the other hand, we have found no support for the idea that Perspective-taking and Meaningful feedback have larger effect on explicit attitudes compared to the Stereotyping mechanic (H6).

These results suggest that particular game design patterns linking depicted topics with some characteristics may differ in their effects on implicit attitudes. However, our findings regarding explicit attitudes do not support our original prediction. There are two possible explanations. First, similar to intervention duration, explicit attitudes may be more prone to being affected by participants' relationships to the measured concept (Kampf, 2016b; Alhabash & Wise, 2015). Second, the chosen category Meaningful feedback is relatively broad. However, that is speculation and further research needs to be done to examine it.

On a general level, our findings about the persuasive game mechanics are the first such complex data brought to the debate about which particular game elements are responsible for attitude change. Focusing on game persuasive mechanics, we have identified which game design patterns are associated with changes in implicit attitudes. However, there was a lack of such clear signals in the domain of explicit attitudes. This particular area of research is still in its beginnings, but our findings are not only relevant for mapping the effect of narrative games on society, they also possess unique value for game designers and experts in education. They can help the latter groups effectively develop games for change. Nevertheless, more studies are needed to fully understand the particular game mechanisms that are affecting our attitudes.

Action vs Non-Action Games

Our analysis did not prove different effect of narrative for action (H7) and non-action games (H8) on implicit or explicit attitudes, so neither hypothesis was supported. These findings suggest that, whether or not the game put players in time pressure situations with a need for fast

reactions and a focus on hand-eye coordination, it has no noticeable effect on players' attitudes. Therefore, the distinction between action and non-action games does not seem to be key in research on attitude change.

Comparator Groups

Available evidence does not indicate that magnitude of explicit attitude change is greater for control groups using topic-unrelated activities than for those using topic-related activities (H3). We assume that the reason games are more effective in attitude change than these more traditional formats is that narrative video games require players to interact with the depicted topics in game narratives in order to proceed further in the gameplay; the games often offer players opportunities to react to the depicted topics.

We have not identified any study allowing us to evaluate the hypothesis related to control groups and implicit attitudes (H4). Therefore, we were not able to test it.

Exploratory Goals: Gender and Age Effects

In our exploratory analysis, we have not identified any effect of gender on the magnitude of attitude change. However, the data suggests that the potential of video games to affect attitudes slightly decreases with age. Again, these outcomes are of an exploratory nature and should be approached with caution. A study by Wan and Chen (2006) suggests that working memory could have a mediating influence on the effect of age on attitude change. Specifically, that attitude change among older adults ($M_{age} = 74.97$) relies more on argument quantity than argument quality compared to younger adults ($M_{age} = 20.03$). The authors suggest this is caused by the limits of working memory at a higher age. However, the weighted mean age across the studies included in the present meta-analysis was relatively low ($M_{age} = 21.13$), thus the suggested limits of working memory are unlikely to be a reason for the difference. Also, several studies disprove the general effect of age on susceptibility to change in attitude (e.g., Tyler & Schuller, 1991;

Krosnick, & Alwin 1989). Accordingly, we see two other variables possibly responsible for this age effect. First, video games as a format may be a less accessible or trustworthy format for older generations. Second, the relationship of older players to the depicted topics in games might be different. All these interpretations require more data for further clarification.

Future Directions

In general, there was sufficient information in the literature to support the main hypothesis. However, the field is still relatively fragmented and more studies are needed to understand fully the effect of video games on players and society. For instance, very few studies collected data on the long-term effects of video games. Also, more studies are focused on explicit attitudes than are on implicit attitudes. Furthermore, research on particular game elements responsible for attitude change is only in its beginnings, so there are no other studies to compare our results with since our analysis of the persuasion mechanics is the first of its kind. Plus, the experiments in the studies included were predominantly focused on the effects of relatively short interventions compared to often multiple-hour-long gameplays offered by most popular video game titles. In relation to this, as our data suggests, the effect of intervention duration on implicit attitudes is worth monitoring. Upcoming studies should explore the effects of longer exposure to video games (for ecological validity reasons). Finally, much more focus, especially when studying explicit attitudes, should be given to players' stances on the topics assessed. This should be mapped using attitude measurements and studying the role of individual characteristics in general. For instance, our exploratory analysis suggests that older participants might be less susceptible to attitude change than younger participants.

Limitations

There are several limitations to this study that are worth mentioning. First, despite the relatively elaborate and labor intensive academic database search for all relevant papers, our final

research sample consists of only 35 papers meeting our criteria. Second, our meta-analysis did not focus in detail on players' characteristics: especially their relationships to depicted topics. That is because the vast majority of studies do not report this data. In our opinion, none of the limitations undermines the paper's key findings.

Conclusion

This meta-analysis is the first contribution of its kind to the body of knowledge in the field of narrative video games and their effects on attitudes. The findings suggest that narrative video games are able to affect players' attitudes towards the topics depicted in these games. This effect is present in studies focused on both implicit and explicit changes in attitude.

Furthermore, our data suggests several moderating effects. Most notably, it seems that longer intervention duration and persuasive game mechanics such as Stereotyping and Meaningful feedback result in larger implicit attitude change. Also, our exploratory analysis suggests that participants' attitudes are less affected by video games as they reach a higher age. However, this effect is rather small.

Narrative video games are widespread phenomena in our societies and culture. Our meta-analysis provides evidence that they are not only part of our world, but that they are also shaping how we think about it. There are still significant gaps in fully understanding this process, and more studies are needed to provide a better picture of how to design video games promoting positive attitudinal change.

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Table 1*Expected Effect of Persuasive Game Mechanics on Explicit and Implicit Attitudes*

Persuasive mechanic	Effect on explicit attitudes	Effect on implicit attitudes
Perspective-taking	significant effect	no significant effect
Meaningful feedback	significant effect	significant effect
Stereotyping	no significant effect	significant effect

Table 2*Meta-analysis Results for Video Games' Effect on Explicit and Implicit Attitudes*

	<i>k</i>	<i>g</i> [95% CI]	<i>SE</i>	τ	<i>I</i> ²	3PSM estimate	3PSM <i>p</i> -value	V&W estimate
Explicit	101	0.24 [0.11, 0.37]	0.07	.43	94%	0.32 [0.09, 0.55]	.006	0.12 [-0.01, 0.26]
Implicit	18	0.36 [0.24, 0.48]	0.05	.11	31%	0.37 [0.20, 0.54]	< .001	0.31 [0.16, 0.46]

Note. Values in brackets represent 95% CI. V&W = Veeva & Woods step function model assuming moderate selection.

Table 3*Substantive Moderators*

Hypothesis	Moderators	Groups of effects	<i>K</i>	Effect size (95% CI)	Statistical test
H2	Duration of intervention		115	$B = 0.05 (-0.03, 0.14)$	$t = 1.31, p = .19$
H3	Control group type & explicit attitude change	Topic-related activities	30	$g = 0.13 (-0.01, 0.28)$	$F(1, 3.67) = 3.05, p = .16$
		Topic-unrelated activities	89	$g = 0.29 (0.14, 0.44)$	
H5	Persuasive mechanics in implicit attitudes	Stereotyping & Meaningful feedback	14	$g = 0.37 (0.24, 0.51)$	$F(1, 6.25) = 10.7, p = .02^*$
		Perspective-taking	2	$g = 0.17 (-0.17, 0.52)$	
H6	Persuasive mechanics in explicit attitudes	Perspective-taking & Meaningful feedback	92	$g = 0.23 (0.08, 0.38)$	$F(1, 3.59) = 2.36, p = .21^*$
		Stereotyping	8	$g = 0.42 (0.10, 0.73)$	
H7 & H8	Type of game	Action games → Implicit attitudes	13	$g = 0.32 (0.21, 0.42)$	$F(1, 3.18) = 1.38, p = .32$
		Action games → Explicit attitudes	31	$g = 0.23 (0.03, 0.42)$	
		Non-action games → Explicit attitudes	69	$g = 0.26 (0.10, 0.42)$	$F(1, 14.30) = 0.18, p = .68$
		Action games → Explicit attitudes	31	$g = 0.21 (0.02, 0.39)$	
Exploratory	Age		91	$B = -0.01 (-0.03, -0.00)$	$t = -2.06, p = .04$
Exploratory	Gender		105	$B = 0.60 (-0.14, 1.33)$	$t = 1.59, p = .11$

Note: * = Satterthwaite small-sample correction used to compute the test and CIs.

Figure 1*Prisma Flow Diagram (Page et al., 2021)*

Figure 2

Forest Plot and Funnel Plot for Video Games' Overall Effect on Attitude Change. In the Forest Plot, Effects are Sorted by Ascending SE

Figure 3*Risk of Bias Chart*

Figure 4*p*-curve or the Overall Effect

Note: The observed *p*-curve includes 4 statistically significant ($p < .05$) results, of which 3 are $p < .025$. There were 9 additional results entered but excluded from *p*-curve because they were $p > .05$.

Appendix A. List of search strings

Table A.1

Search strings used to search the academic databases

Database	Search string	Number of articles 18. 8. 2020	Note
Scopus	TITLE-ABS-KEY((attitude* OR stereotype*) AND (game*) AND (change OR effect OR significant* OR impact) AND (experiment* OR empirical* OR intervention))	1464	Searched in Title, Abstract, and Keywords
The ACM Guide to Computing Literature	[[Abstract: attitude*] OR [Abstract: stereotype*]] AND [Abstract: game*] AND [[Abstract: change] OR [Abstract: effect] OR [Abstract: significant*] OR [Abstract: impact]] AND [[Abstract: experiment*] OR [Abstract: empirical*] OR [Abstract: intervention]]	125	Searched in abstract only due to the limitation of a platform
Eric	(title:(attitude attitudes stereotype stereotypes) OR abstract:(attitude attitudes stereotype stereotypes)) AND (title:(game games) OR abstract:(game games)) AND (title:(change effect effects significant significantly impact) OR abstract:(change effect effects significant significantly impact)) AND (title:(experiment experimental experimentally empirical empirically intervention) OR abstract:(experiment experimental experimentally empirical empirically intervention))	174	Searched in Title and Abstract
Science direct*	(attitude OR stereotype) AND (game) AND (change OR effect OR significant OR impact) AND (experiment OR experimental)	100	Searched in Title, Abstract, and Keywords
Science direct*	(attitude OR stereotype) AND (game) AND (change OR effect OR significant OR impact) AND (experimentally OR empirical)	84	Searched in Title, Abstract, and Keywords
Science direct*	(attitude OR stereotype) AND (game) AND (change OR effect OR significant OR impact) AND (empirically OR intervention)	65	Searched in Title, Abstract, and Keywords
Web of Science Core Collection	(attitude* OR stereotype*) AND (game*) AND (change OR effect OR significant* OR impact) AND (experiment* OR empirical* OR intervention)	1170	Searched in Topic, Title, Abstract, and Keywords
ProQuest	ti,ab(attitude* OR stereotype*) AND ti,ab(game*) AND ti,ab(change OR effect OR significant* OR impact) AND ti,ab(experiment* OR empirical* OR intervention)	650	Searched in Title and Abstract
Google Scholar*	attitude* AND game*	1000	The identified studies from full text search during 1. 9 2018 were

included into the
sample

Note. *At the time of the search, Science direct did not support wildcards and had limited number of characters for one search operation, therefore, we have divided our search operator into the three

Appendix B. The effect of excluding the keyword “significant”.

The exclusion of the word “significant” from the search string resulted in missing two studies, Cangas et al. (2017) reporting two relevant effects and Cicchirillo (2009) reporting one relevant effect. As the latter study did not report the direction of the effect, it was not included in any meta-analytic models anyway. The exclusion of the two effects from Cangas (both were deemed relevant and included in our models) lead to an drop in our main estimate (H1) by $g = .0051$. Thus, the exclusion of the keyword “significant” can be considered inconsequential with respect to the meta-analytic set of effects or the actual results.

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Appendix E. Exploratory bias-adjustment methods

As a secondary approach to publication bias adjustment, we employed the regression-based PET-PEESE method. This model regresses the effect size on a measure of precision. Given that the regression slope reflects the presence of small-study effects (as larger studies are less

likely to stay unpublished), the model intercept can then be interpreted as an average ES for a hypothetical, infinitely precise study (Stanley & Doucouliagos, 2014). In line with our previous work (see IJzerman et al., 2021), we used a multi-level RVE-based implementation of PET-PEESE, modeling the same hierarchical structure as in our random-effects meta-analytic models. Apart from that, we used $\sqrt{(2/N)}$ and a $2/N$ terms instead of standard error and variance for PET and PEESE, respectively, as an estimate of precision. That is because N -based predictors do not induce a correlation between effect size and variance (as the latter is calculated using the former), and models fitted using N -based predictors exhibit a markedly lower false-positive rate (Pustejovsky, 2017). As a sensitivity analysis, we also employed the selection model as a conditional estimator for PET-PEESE instead of the traditional PET, since the former tends to fare better in terms of precision and error rates in similar analytic scenarios (Carter et al. 2019). Dependencies among the effects were directly modeled using a RVE-based multilevel regression model with the exact same random effects structure as the naive meta-analytic model. The main results of the PET-PEESE analyses are reported in the paper.

In the supplementary analysis output, we also report the results of another two bias-adjustment methods, the p -uniform* (van Aert & van Assen, 2021) and Weighted Average of the Adequately Powered studies (WAAP-WLS; Stanley et al., 2016). The effect size estimation by p -uniform* makes use of the fact that p -values follow a uniform distribution at the true effect size. The combined WAAP-WLS estimator tries to identify studies that are adequately powered to detect the meta-analytic effect. If there are fewer than two such studies, the method falls back to the WLS estimator (Stanley, Doucouliagos, & Ioannidis, 2017). If there are at least two adequately powered studies, WAAP returns a WLS estimate based on effects from only those studies.

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